

Displays for Statistics 5303

Lecture 28

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Class Web Page

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Skeleton ANOVA tables are important for testing and estimation.

One factor skeleton table

Source	DF	EMS
Treatments	a-1	$\sigma^2 + n\sigma_{\alpha}^2$
Error	N-a	$\sigma^2 = \sigma_{\epsilon}^2$

Two factor skeleton table

Source	DF	EMS
A	a-1	$\sigma^2 + n\sigma_{\alpha\beta}^2 + nb\sigma_{\alpha}^2$
B	b-1	$\sigma^2 + n\sigma_{\alpha\beta}^2 + na\sigma_{\beta}^2$
AB	(a-1)(b-1)	$\sigma^2 + n\sigma_{\alpha\beta}^2$
Error	ab(n-1)	σ^2

The multiplier of a term is the number of cases affected by one effect of that type

- 1 case is affected by each ϵ_{ijk}
- n cases are affected by each $\alpha\beta_{ij}$
- nb cases are affected by each α_i
- na cases are affected by each β_j

For the box-making machines, a = 10, b = 10, n = 4 so the table is

Source	DF	EMS
A:Machines	9	$\sigma^2 + 4\sigma_{\alpha\beta}^2 + 40\sigma_{\alpha}^2$
B:Operators	9	$\sigma^2 + 4\sigma_{\alpha\beta}^2 + 40\sigma_{\beta}^2$
AB	81	$\sigma^2 + 4\sigma_{\alpha\beta}^2$
Error	300	σ^2

Note that $EMS_A = EMS_{AB} + 40\sigma_{\alpha}^2$.

This means that $EMS_A = EMS_{AB}$ if and only if $\sigma_{\alpha}^2 = 0$.

$F = MS_1/MS_2$ really tests $H_0: E(MS_1) = E(MS_2)$. So to test $H_0: \sigma_{\alpha}^2 = 0$ the proper F-statistic is $F = MS_A/MS_{AB}$ (denominator = MS_{AB}).

This is *different* from the fixed effect case where you use $F = MS_A/MS_E$ (denominator = MS_{error}) to test $H_0: \text{all } \alpha_i = 0$.

Three factor skeleton table

Source	DF	EMS
A	a-1	$\sigma^2 + n\sigma_{\alpha\beta\gamma}^2 + nc\sigma_{\alpha\beta}^2 + nb\sigma_{\alpha\gamma}^2 + nbc\sigma_{\alpha}^2$
B	b-1	$\sigma^2 + n\sigma_{\alpha\beta\gamma}^2 + nc\sigma_{\alpha\beta}^2 + na\sigma_{\beta\gamma}^2 + nac\sigma_{\beta}^2$
C	c-1	$\sigma^2 + n\sigma_{\alpha\beta\gamma}^2 + nb\sigma_{\alpha\gamma}^2 + na\sigma_{\beta\gamma}^2 + nab\sigma_{\gamma}^2$
AB	(a-1)(b-1)	$\sigma^2 + n\sigma_{\alpha\beta\gamma}^2 + nc\sigma_{\alpha\beta}^2$
AC	(a-1)(c-1)	$\sigma^2 + n\sigma_{\alpha\beta\gamma}^2 + nb\sigma_{\alpha\gamma}^2$
BC	(b-1)(c-1)	$\sigma^2 + n\sigma_{\alpha\beta\gamma}^2 + na\sigma_{\beta\gamma}^2$
ABC	(a-b)(b-1)(c-1)	$\sigma^2 + n\sigma_{\alpha\beta\gamma}^2$
Error	abc(n-1)	σ^2

- n cases affected by each $\alpha\beta\gamma_{ijk}$
- nc cases affected by each $\alpha\beta_{ij}$
- nbc cases affected by each α_i , etc.

Here's a an example of a balanced one factor random effect experiment. The data are weights of calves sired by a = 5 bulls, n = 8 calves per bull

```
Cmd> sire <- factor(1,1,1,1,1,1,1, 2,2,2,2,2,2,2,\
3,3,3,3,3,3,3, 4,4,4,4,4,4,4, 5,5,5,5,5,5,5)
Cmd> wts <- vector(61,100,56,113,99,103,75,62,\
75,102,95,103,98,115,98,94, 58,60,60,57,57,59,54,100,\
57,56,67,59,58,121,101,101, 59,46,120,115,115,93,105,75)
Cmd> anova("wts=sire",fstat:T)
Model used is wts=sire
      DF      SS      MS      F      P-value
CONSTANT  1  2.7258e+05  2.7258e+05  587.71949  0
sire      4    5591.1    1397.8    3.01382  0.030874
ERROR1   35    16233    463.79
```

The interest here is the contribution to the variability of weights due to parent.

ems() computes EMS formulas

```
Cmd> ems("wts=sire","sire")
EMS(CONSTANT) = V(ERROR1) + 8V(sire) + 40Q(CONSTANT)
EMS(sire) = V(ERROR1) + 8V(sire)
EMS(ERROR1) = V(ERROR1)
```

V(ERROR1) stands for σ^2 .

V(sire) stands for σ_α^2

Q(CONSTANT) stands for μ^2 , a function of the fixed parameter μ

When data are unbalanced, the formulas are harder but can be computed by ems().

Here I set 4 reponses to MISSING and ran ems() again.

```
Cmd> wts1 <- wts; wts1[vector(2, 11, 12, 29,30)] <- ?
Cmd> tabs(wts1,sire,count:T) # it's now unbalanced
WARNING: MISSING values in argument 1 to tabs() omitted
(1)      7      6      8      6      8
Cmd> ems("wts1=sire","sire")
EMS(CONSTANT) = V(ERROR1) + 7.1143V(sire) + 35Q(CONSTANT)
EMS(sire) = V(ERROR1) + 6.9714V(sire)
EMS(ERROR1) = V(ERROR1)
```

$$EMS_{sire} = \sigma^2 + 6.9714\sigma_\alpha^2$$

This tells you that $F = MS_A / MS_{error}$ is still OK for testing $H_0: \sigma_\alpha^2 = 0$.

But $F = MS_{const} / MS_A$ is no longer OK to test $\mu = 0$, since

$$EMS_{constant} - EMS_A = 35\mu^2 + 0.1429\sigma_A^2$$

From the output

$$EMS_{constant} = \sigma^2 + 8\sigma_\alpha^2 + 40\mu^2$$

$$EMS_A = \sigma^2 + 8\sigma_\alpha^2$$

The multipliers here are $n = 8$ and $n \times a = 40$.

If there was any reason to test $H_0: \mu = 0$ (there isn't in this case), the formulas show you that the proper F-statistic would be $F = MS_{constant} / MS_A$ on 1 and 4 d.f.

```
Cmd> 2.7258e+05/1397.8
(1) 195.01
Cmd> 1 - cumF(195.01,1,4)
(1) 0.00015252
```

Once you get beyond two-way designs, testing gets more complicated.

Suppose you want to test $H_0: \sigma_\alpha^2 = 0$:

$$EMS_A = \sigma^2 + n\sigma_{\alpha\beta\gamma}^2 + nc\sigma_{\alpha\beta}^2 + nb\sigma_{\alpha\gamma}^2 + nbc\sigma_\alpha^2$$

When H_0 is true,

$$EMS_A = \sigma^2 + n\sigma_{\alpha\beta\gamma}^2 + nc\sigma_{\alpha\beta}^2 + nb\sigma_{\alpha\gamma}^2$$

but there is no term with this EMS to use as a denominator MS in an F-statistic.

You need to find a numerator and denominator MS such that

$$E(MS_{num}) - E(MS_{den}) = const \times \sigma_\alpha^2$$

so that when you compare MS_{num} and MS_{den} using $F = MS_{num} / MS_{den}$ you are comparing two quantities whose means are the same when H_0 is true.

One approach (not a good one, but a natural one):

Include both MS_{AB} and MS_{AC} in the denominator so that EMS contains both $n\sigma_{\alpha\beta}^2$ and $nb\sigma_{\alpha\gamma}^2$. Since the EMS now includes $2 \times n\sigma_{\alpha\beta\gamma}^2$ and EMS_A has only $n\sigma_{\alpha\beta\gamma}^2$, also subtract MS_{ABC} to get rid of the extra $n\sigma_{\alpha\beta\gamma}^2$. This leads to

$$F = MS_A / (MS_{AB} + MS_{AC} - MS_{ABC}).$$

- **Advantage:** $MS_{num} = MS_A$, the fixed effects numerator.
- **Disadvantage:** It's possible to have $MS_{den} < 0$ and hence $F < 0$ which can never happen with a real F-statistic.

The better approach is to find MS_{num} and MS_{den} using only positive coefficients.

Here is an analysis of the data used but not listed in Oehlert Example 11.2. It is artificial data purporting to be measurements of carton strength.

```

Cmd> carton3 <- read( "", "carton3" )
carton3      400      4
Read from file "TP1:Stat5303:Data:carton.dat"

Cmd> makecols( carton3, mach, oper, gbat, y )

Cmd> mach <- factor( mach ); oper <- factor( oper )
Cmd> gbat <- factor( gbat ) # glue batch

Cmd> anova( "y=mach*oper*gbat", pval:T )
Model used is y=mach*oper*gbat

```

	DF	SS	MS	P-value
CONSTANT	1	8.6671e+06	8.6671e+06	0
mach	9	2705.8	300.64	3.4897e-16
oper	9	8886.8	987.42	8.0281e-42
mach.oper	81	1682.5	20.772	0.71494
gbat	1	2375.8	2375.8	1.1082e-19
mach.gbat	9	420.48	46.72	0.039738
oper.gbat	9	145.34	16.149	0.71282
mach.oper.gbat	81	1649.8	20.368	0.74902
ERROR1	200	4645.8	23.229	

```

Cmd> ems( "y=mach*oper*gbat", vector( "mach", "oper", "gbat" ) )
Compacting memory, please stand by in macro colproduct
EMS( CONSTANT ) = V( ERROR1 ) + 2V( mach.oper.gbat ) + 20V( oper.gbat )
+ 20V( mach.gbat ) + 200V( gbat ) + 4V( mach.oper ) + 40V( oper ) +
40V( mach ) + 400Q( CONSTANT )
EMS( mach ) = V( ERROR1 ) + 2V( mach.oper.gbat ) + 20V( mach.gbat ) +
4V( mach.oper ) + 40V( mach )
EMS( oper ) = V( ERROR1 ) + 2V( mach.oper.gbat ) + 20V( oper.gbat ) +
4V( mach.oper ) + 40V( oper )
EMS( mach.oper ) = V( ERROR1 ) + 2V( mach.oper.gbat ) + 4V( mach.oper )
EMS( gbat ) = V( ERROR1 ) + 2V( mach.oper.gbat ) + 20V( oper.gbat ) +
20V( mach.gbat ) + 200V( gbat )
EMS( mach.gbat ) = V( ERROR1 ) + 2V( mach.oper.gbat ) + 20V( mach.gbat )
EMS( oper.gbat ) = V( ERROR1 ) + 2V( mach.oper.gbat ) + 20V( oper.gbat )
EMS( mach.oper.gbat ) = V( ERROR1 ) + 2V( mach.oper.gbat )
EMS( ERROR1 ) = V( ERROR1 )

```

Approach using positive coefficients

Include MA_{ABC} in MS_{num} to compensate for the extra $n\sigma_{\alpha\beta}^2$ in $E(MS_A + MS_{ABC})$.

$$F = (MS_A + MS_{ABC}) / (MS_{AB} + MS_{AC})$$

$$\begin{aligned}
 E(MS_{denom}) &= E(MS_{AB} + MS_{AC}) \\
 &= 2\sigma^2 + 2n\sigma_{\alpha\beta\gamma}^2 + n\sigma_{\alpha\beta}^2 + nb\sigma_{\alpha\gamma}^2 \\
 E(MS_{num}) &= E(MS_A + MS_{ABC}) \\
 &= 2\sigma^2 + 2n\sigma_{\alpha\beta\gamma}^2 + n\sigma_{\alpha\beta}^2 + nb\sigma_{\alpha\gamma}^2 + nbc\sigma_{\alpha}^2 \\
 &= E(MS_{denom}) + nbc\sigma_{\alpha}^2
 \end{aligned}$$

Unfortunately, when H_0 is true, F does not have an F-distribution, although an F-distribution with specially computed degrees of freedom provides a pretty good approximation.

As you can see, my Mac complained about the need for lots of memory to compute the EMS table.

You can check the coefficients match the formulas. For instance $n = 2$ is always the multiplier for $V(mach.oper.gbat) = \sigma_{\alpha\beta\gamma}^2$ and $nac = 40$ is the multiplier for $V(oper) = \sigma_{\beta}^2$

Compute the F-statistics to test

$$H_0: \sigma_{\alpha}^2 = 0$$

```

Cmd> ms_num <- MS[2] + MS[8]
Cmd> ms_denom <- MS[4] + MS[6]
Cmd> f_stat <- ms_num/ms_denom; f_stat
(1) 4.7563

```

Q. Since F doesn't really have the F-distribution, how do you use it to test H_0 ?

A. You still use the F-distribution, but with special calculations for degrees of freedom, as an approximation to the distribution when H_0 is true

In this case, the formulas for the degrees of freedom are.

$$df_{num} = \frac{(MS_A + MS_{ABC})^2}{MS_A^2/df_A + MS_{ABC}^2/df_{ABC}}$$

$$= MS_{num}^2 / \{MS_A^2/df_A + MS_{ABC}^2/df_{ABC}\}$$

$$df_{denom} = \frac{(MS_{AB} + MS_{AC})^2}{MS_{AB}^2/df_{AB} + MS_{AC}^2/df_{AC}}$$

$$= MS_{denom}^2 / \{MS_{AB}^2/df_{AB} + MS_{AC}^2/df_{AC}\}$$

This approximation is due to Satterthwaite.

Estimates of variance components

There are several ways to estimate variance components.

Simplest and easiest to understand:

Use a **linear combination of MS** that has the proper expectation.

For the one-way balanced case

$$EMS_A = \sigma^2 + n\sigma_{\alpha}^2 \text{ and } EMS_{error} = \sigma^2$$

so $(EMS_A - EMS_{error})/n = \sigma_{\alpha}^2$ and

$$\hat{\sigma}_{\alpha}^2 = (MS_A - MS_{error})/n \text{ is unbiased}$$

For the two-way balanced case:

$$EMS_A = \sigma^2 + n\sigma_{\alpha\beta}^2 + nb\sigma_{\alpha}^2$$

$$EMS_{AB} = \sigma^2 + n\sigma_{\alpha\beta}^2, EMS_{error} = \sigma^2$$

Then $(EMS_{AB} - EMS_{error})/n = \sigma_{\alpha\beta}^2$

$$(EMS_A - EMS_{AB})/(nb) = \sigma_{\alpha}^2$$

So unbiased estimates are

$$\hat{\sigma}_{\alpha\beta}^2 = (MS_{AB} - MS_{ABC})/n$$

$$\hat{\sigma}_{\alpha}^2 = (MS_A - MS_{AB})/(nb)$$

```
Cmd> ms_num <- MS[2] + MS[8]
Cmd> ms_denom <- MS[4] + MS[6]
Cmd> f_stat <- ms_num/ms_denom; f_stat
(1) 4.7563
Cmd> df_num <- ms_num^2/(MS[2]^2/DF[2] + MS[4]^2/DF[4])
Cmd> df_denom <- ms_denom^2/(MS[4]^2/DF[4] + MS[6]^2/DF[6])
Cmd> vector(df_num,df_denom)
(1) 10.255 18.378
Cmd> 1 - cumF(f_stat,df_num,df_denom)
(1) 0.0018512
```

Macro mixed() does this for you automatically:

```
Cmd> mixed("y=mach*oper*gbat",vector("mach","oper","gbat"))
      DF      MS      Error DF      Error MS      F      P value
CONSTANT      1 8.667e+06      2.355      3684      2353 0.0001374
mach      10.26      321      18.38      67.49      4.756 0.001851
oper      9.375      1008      39.74      36.92      27.3 1.765e-14
mach.oper      81      20.77      81      20.37      1.02 0.4648
gbat      1.017      2396      14.56      62.87      38.11 1.915e-05
mach.gbat      9      46.72      81      20.37      2.294 0.02386
oper.gbat      9      16.15      81      20.37      0.7929 0.6237
mach.oper.gbat 81      20.37      200      23.23      0.8768 0.749
ERROR1      200      23.23      0      0      MISSING MISSING
```

Here's the **general formula for DF**.

When $MS = \sum_k g_k MS_k$, where MS_k has df_k degrees of freedom, approximately

$$DF = MS^2 / (\sum_k g_k^2 MS_k^2 / df_k)$$

When all the $g_k = 1$, $DF = MS^2 / (\sum_k MS_k^2 / df_k)$

For the three-way balanced case, since $EMS_A - EMS_{AB} - EMS_{AC} + EMS_{ABC} = nb\sigma_{\alpha}^2$

$$\hat{\sigma}_{\alpha}^2 = (MS_A - MS_{AB} - MS_{AC} + MS_{ABC})/nb$$

is unbiased.

```
Cmd> (MS[2]-MS[4]-MS[6]+MS[8])/(2*2*10)
(1) 6.338
```

You can calculate approximate degrees of freedom similarly as before as $df =$

$$\frac{(MS_A - MS_{AB} - MS_{AC} - MS_{ABC})^2}{MS_A^2/df_A + MS_{AB}^2/df_{AB} + MS_{AC}^2/df_{AC} + MS_{ABC}^2/df_{ABC}}$$

```
Cmd> J <- vector(2,4,6,8)
```

```
Cmd> (MS[2]-MS[4]-MS[6]+MS[8])^2/sum(MS[J]^2/DF[J])
(1) 6.2425
```

varcomp() does black box computations.

```
Cmd> varcomp("y=mach*oper*gbat",vector("mach","oper","gbat"))
      Estimate      SE      DF
mach      6.338      3.5875      6.2425
oper      24.272      11.639      8.6976
mach.oper 0.10114      1.1428      0.015664
gbat      11.666      16.8      0.96449
mach.gbat 1.3176      1.1128      2.8042
oper.gbat -0.21093      0.41291      0.52191
mach.oper.gbat -1.4307      1.9773      1.0471
ERROR1      23.229      2.3229      200
```

SE is almost meaningless here because sample sizes are very small.