Statistics 5301 — Final Exam Sketched Solutions — June 8, 1994

1. a) We have a 2^3 factorial treatment structure. Since recipe is difficult to change, it might make sense to do a split plot experiment with batch as whole plot and recipe as whole plot treatment, and the four mold/cooling combinations as split plot treatments. All effects are fixed and crossed; whole plots design is completely randomized.

b) The units are the 9 groups of 400 voters, the treatments are the three bits of information (fixed). We can do a completely randomized assignment of the treatments to the units.

c) Rose bush is unit; volunteer and moisture are extraneous sources of variation to block. We could do a replicated 4 by 4 Latin square, where moisture (row) and volunteer (column) are the blocking variables and we have 4 fungicide treatments. All treatments are fixed.

d) We have a 2^4 factorial treatment structure (16 combinations, factors fixed and crossed) but can only afford 8 units. Thus a fractional factorial is appropriate. Use I=ABCD as a generator.

e) Feet are units and subject is block. Since people only have two feet and we have four treatments (the four powders, fixed), we have incomplete blocks. There are 6 pairs of treatments, and since we have 60 boys, we can use each pair 10 times for a BIBD design. Another possible alternative would be a crossover design, though you might have to balance for residual effects.

f) We have a 2^4 factorial treatment structure with all factors fixed and crossed. The cooler must serve as block, however, and it is of size 8. Thus we need incomplete blocks, and confounding is suggested. We could partially confound, using ABCD in replicate 1 and BCD in replicate 2.

- a) This is a completely randomized design with children as unit and five fixed treatments. Cavities on nonmolars is a covariate. The treatment structure is a little funky, being a 2 by 2 factorial plus an additional control.
- b) This is a fully nested design, with brand (fixed) on top, day (random) nested in brand, and carton (random) nested in day.
- c) This is a split plot, with branch as whole plot (blocked into high and low sales blocks), hardware as whole plot treatment, week as split plot, and software as split plot treatment. Factors are fixed and crossed.

Source	aı						
cov	1	_					
trt	4						
error	34						
Source		df					
brand		1					
day (brand)							
$\operatorname{carton}(\operatorname{day})$							
error		30					
Source	df						
block	1	-					
hardware	1						
wp error	17						
software	2	_					
hard.soft	2						
sp error	36						

10

Source		df
block		3
glue		2
curing		1
glue.curi	ng	2
error		15
Source	df	
block	5	
variety	5	
error	13	
Source	df	_
clearing	1	_
hour	2	
\mathbf{phase}	2	
trt	2	
error	10	

- d) This is a randomized complete block design with batch as block. The treatments have a 2 by 3 factorial structure and are crossed and fixed.
- e) This is a partially balanced incomplete block. For example, A occurs 4 times with D, but only twice with B, C, E, and F. Treatments are fixed.
- f) This is a replicated Latin square. Treatments are the three recorded calls (fixed), and the two blocking factors are phase of the breeding season and morning hour. The two clearings are the two replicates.

3. Start with a full factorial in A, B, C, D and add on E and F. -ACD ABC

					-ACD	ADU				
	Α	В	С	D	=E	=F	Ι	-ACDE	ABCF	-BDEF
e	-	-	-	-	+	-	A	-CDE	BCF	-ABDEF
af	+	-	-	-	-	+	В	-ABCDE	ACF	-DEF
\mathbf{bef}	-	+	-	-	+	+	AB	-BCDE	CF	-ADEF
$^{\mathrm{ab}}$	+	+	-	-	-	-	\mathbf{C}	-ADE	ABF	-BCDEF
\mathbf{cf}	-	-	+	-	-	+	\mathbf{AC}	-DE	$_{\mathrm{BF}}$	-ABCDEF
ace	+	-	+	-	+	-	BC	-ABDE	\mathbf{AF}	-CDEF
\mathbf{bc}	-	+	+	-	-	-	ABC	-BDE	\mathbf{F}	-ACDEF
abcef	+	+	+	-	+	+	D	-ACE	ABCDF	-BEF
d	-	-	-	+	-	-	AD	-CE	BCDF	-ABEF
adef	+	-	-	+	+	+	BD	-ABCE	ACDF	$-\mathrm{EF}$
bdf	-	+	-	+	-	+	ABD	-BCE	CDF	-AEF
abde	+	+	-	+	+	-	CD	-AE	ABDF	-BCEF
cdef	-	-	+	+	+	+	ACD	-E	BDF	-ABCEF
acd	+	-	+	+	-	-	BCD	-ABE	ADF	-CEF
bcde	-	+	+	+	+	-	ABCD	-BE	DF	-ACEF
abcdf	+	+	+	+	-	+				

4. Blocks (α_i) are random, and meters random and nested in block $(\beta_{j(i)})$. Weeks are random (γ_k) . I would say that day (δ_l) is fixed and crossed with the other effects, because day of week effects are likely to be similar from week to week and block to block. Weeks, day, and meter are crossed. All interactions are random.

$$y_{ijkl} = \alpha_i + \beta_{j(i)} + \gamma_k + \alpha \gamma_{ik} + \beta \gamma_{jk(i)} + \delta_l + \alpha \delta_{il} + \beta \delta_{jl(i)} + \gamma \delta_{kl} + \alpha \gamma \delta_{ikl} + \beta \gamma \delta_{jkl(i)}$$

b) The skeleton ANOVA table also includes the expected mean squares and tests to use, though

	Source	df	EMS	5			test
	α_j	7	$7\sigma_{\beta\gamma}^2$	$1+35\sigma_{lpha}^2$	$_{\gamma} + 42\sigma_{\beta}^2 + 210\sigma_{\alpha}^2$		(A+BC(A))/(B(A)+AC)
	$\beta_{j(i)}$	32	$7\sigma_{\beta\gamma}^2$	$+42\sigma_{\beta}^2$, , , , , , , , , , , , , , , , , , ,		B(A)/BC(A)
	γ_k	5	$7\sigma_{\beta\gamma}^2$	$+35\sigma_{lpha}^2$	$_{\gamma}+280\sigma_{\gamma}^{2}$		C/AC
	$lpha\gamma_{ik}$	35	$7\sigma_{\beta\gamma}^2$	$+35\sigma_{lpha}^2$	γ		AC/BC(A)
	$\beta \gamma_{jk(i)}$	160	$7\sigma_{\beta\gamma}^2$				none
	δ_l	6	$\sigma^2_{\beta\gamma\delta}$	$+5\sigma_{\alpha\gamma\delta}^2$	$+40\sigma_{\gamma\delta}^2+6\sigma_{\beta\delta}^2+$	$+30\sigma_{lpha\delta}^2+240\sum {\delta_l^2}/7$	(D+ACD)/(AD+CD)
	$lpha \delta_{il}$	42	$\sigma^2_{\beta\gamma\delta}$	$+5\sigma_{\alpha\gamma\delta}^{2}$	$+ 6\sigma^2_{\beta\delta} + 30\sigma^2_{\alpha\delta}$		(AD+BCD(A))/(BD(A)+ACD)
	$\beta \delta_{jl(i)}$	192	$\sigma^2_{\beta\gamma\delta}$	$+6\sigma_{\beta\delta}^{2}$	r		BD/BCD(A)
	$\gamma \delta_{kl}$	30	$\sigma^2_{\beta\gamma\delta}$	$+5\sigma_{\alpha\gamma\delta}^2$	$+40\sigma_{\gamma\delta}^2$		CD/ACD
	$lpha\gamma\delta_{ikl}$	210	$\sigma^2_{\beta\gamma\delta}$	$+5\sigma_{\alpha\gamma\delta}^{2}$	1		ACD/BCD(A)
	$eta\gamma\delta_{jkl(i)}$	960	$\sigma^2_{\beta\gamma\delta}$,			none
	Source	df	\mathbf{SS}	MS	F		
5.	trt	4 1	04.69	26.17		The F ratio is 6.52 ,	with 4 and 15 df. This
	error	15	60.18	4.012		gives a p-value of 0 the null hypothesis (.003, so we would reject that all means are equal.

that was not required in the problem.

b) We could use the contrast with coefficients (-3, -3, 2, 2, 2). This contrast has estimated value 7.0852, with SE = $(MSE \sum c_i^2/n_i)^{1/2} = 5.485$, giving a t ratio of 7.0852/5.485 = 1.29. This t-ratio has a two sided p-value of about 0.2, so the null hypothesis is not rejected.