

```
> library(Stat5303libs); library(cfcdae); library(lme4)
```

```
> weardata <- read.table("wear.dat.txt", header=TRUE)
```

I'm embarrassed to say, but I can't really remember where I got these data. I think from one of my professors back in the dark ages. The issue is wear loss of rubber, and we have seven different treatments for reducing the loss. A block is a chunk of rubber, but a block is only large enough for four of the seven treatments. Small losses are good.

```
> weardata
```

| | block | trt | wear |
|----|-------|-----|------|
| 1 | 2 | 1 | 344 |
| 2 | 4 | 1 | 337 |
| 3 | 6 | 1 | 369 |
| 4 | 7 | 1 | 396 |
| 5 | 1 | 2 | 627 |
| 6 | 4 | 2 | 537 |
| 7 | 5 | 2 | 520 |
| 8 | 7 | 2 | 602 |
| 9 | 2 | 3 | 233 |
| 10 | 3 | 3 | 251 |
| 11 | 5 | 3 | 278 |
| 12 | 7 | 3 | 240 |
| 13 | 1 | 4 | 248 |
| 14 | 3 | 4 | 211 |
| 15 | 6 | 4 | 196 |
| 16 | 7 | 4 | 273 |
| 17 | 3 | 5 | 160 |
| 18 | 4 | 5 | 195 |
| 19 | 5 | 5 | 199 |
| 20 | 6 | 5 | 185 |
| 21 | 1 | 6 | 563 |
| 22 | 2 | 6 | 442 |
| 23 | 5 | 6 | 595 |
| 24 | 6 | 6 | 606 |
| 25 | 1 | 7 | 252 |
| 26 | 2 | 7 | 226 |
| 27 | 3 | 7 | 297 |
| 28 | 4 | 7 | 300 |

```
> weardata <- within(weardata, {block <- as.factor(block); trt <- as.factor(trt)})
```

```
> with(weardata, table(block, trt))
```

The incidence matrix n_{ij} . Each block contains four treatments, and each pair of treatments occurs together twice.

| | trt | | | | | | |
|-------|-----|---|---|---|---|---|---|
| block | 1 | 2 | 3 | 4 | 5 | 6 | 7 |
| 1 | 0 | 1 | 0 | 1 | 0 | 1 | 1 |
| 2 | 1 | 0 | 1 | 0 | 0 | 1 | 1 |
| 3 | 0 | 0 | 1 | 1 | 1 | 0 | 1 |
| 4 | 1 | 1 | 0 | 0 | 1 | 0 | 1 |
| 5 | 0 | 1 | 1 | 0 | 1 | 1 | 0 |
| 6 | 1 | 0 | 0 | 1 | 1 | 1 | 0 |
| 7 | 1 | 1 | 1 | 1 | 0 | 0 | 0 |

```
> 7*3/(6*4)
```

Relative efficiency is 7/8.

```
[1] 0.875
```

```
> .875*4
```

Effective sample size for intrablock analysis is 3.5.

```
[1] 3.5
```

```
> fit1 <- lm(wear~block+trt,data=weardata)
```

Basic model is treatments adjusted for blocks. This is the “intrablock” analysis.

Analysis of Variance Table

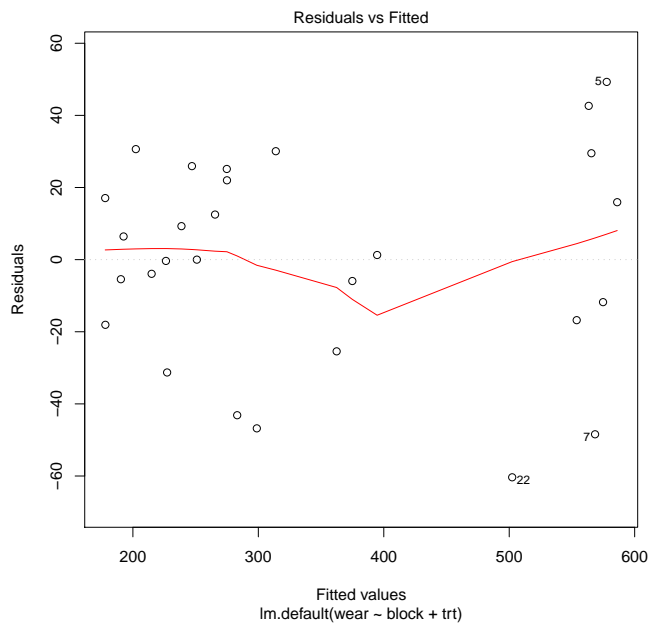
Response: wear

| | Df | Sum Sq | Mean Sq | F value | Pr(>F) | |
|-----------|----|--------|---------|---------|-----------|-----|
| block | 6 | 97395 | 16232 | 11.032 | 8.937e-05 | *** |
| trt | 6 | 506799 | 84466 | 57.404 | 1.687e-09 | *** |
| Residuals | 15 | 22071 | 1471 | | | |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

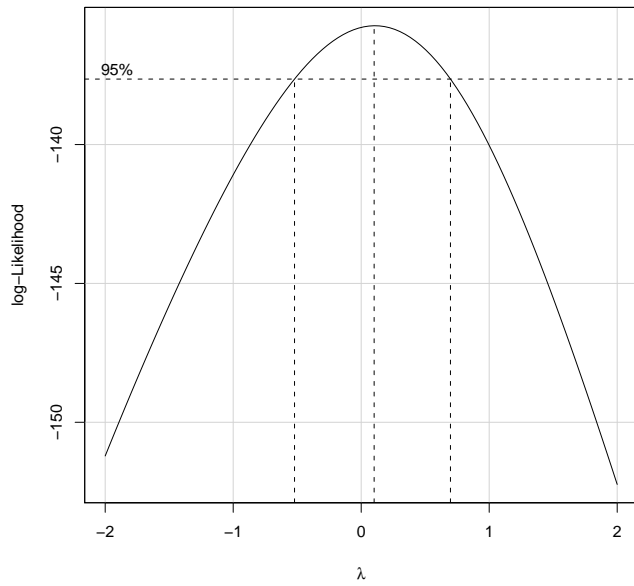
```
> plot(fit1, which=1)
```

Variance could be increasing with mean. Just the slightest touch of curvature.



```
> boxCox(fit1)
```

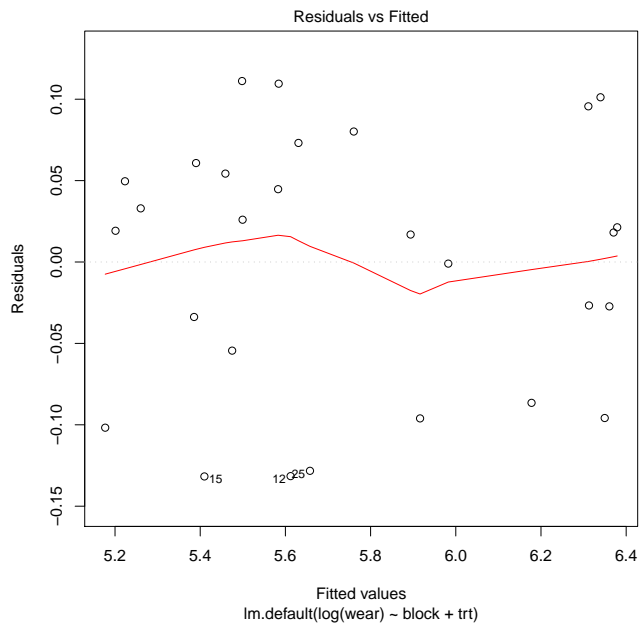
Box Cox suggests a log.



```
> fit2 <- lm(log(wear) ~ block + trt, data = weardata)
```

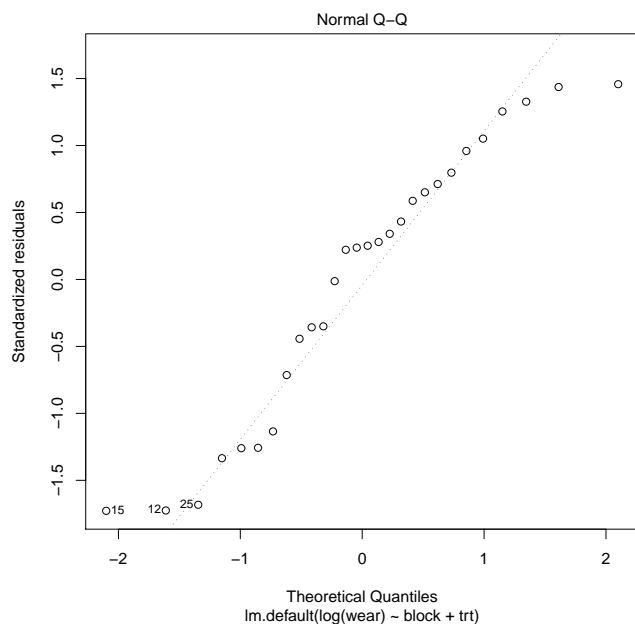
```
> plot(fit2, which = 1)
```

Variance is more constant.



```
> plot (fit2, which=2)
```

Residuals are somewhat short tailed.



```
> anova (fit2)
```

We want treatments after blocks (or use Anova to get type II). Treatments are highly significant.

Analysis of Variance Table

Response: log(wear)

| | Df | Sum Sq | Mean Sq | F value | Pr(>F) | |
|-----------|----|--------|---------|---------|-----------|-----|
| block | 6 | 0.7784 | 0.12974 | 11.959 | 5.573e-05 | *** |
| trt | 6 | 3.9060 | 0.65100 | 60.006 | 1.230e-09 | *** |
| Residuals | 15 | 0.1627 | 0.01085 | | | |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```
> summary (fit2)
```

Call:

```
lm.default(formula = log(wear) ~ block + trt)
```

Residuals:

| Min | 1Q | Median | 3Q | Max |
|----------|----------|---------|---------|---------|
| -0.13172 | -0.06245 | 0.01866 | 0.05593 | 0.11117 |

Coefficients:

| | Estimate | Std. Error | t value | Pr(> t) | |
|-------------|-----------|------------|---------|----------|-----|
| (Intercept) | 5.757176 | 0.019684 | 292.480 | < 2e-16 | *** |
| block1 | 0.040494 | 0.051545 | 0.786 | 0.444326 | |
| block2 | -0.142221 | 0.051545 | -2.759 | 0.014614 | * |
| block3 | -0.033005 | 0.051545 | -0.640 | 0.531628 | |

```

block4      0.013441  0.051545  0.261 0.797822
block5      0.050401  0.051545  0.978 0.343675
block6     -0.008792  0.051545 -0.171 0.866844
trt1        0.145530  0.051545  2.823 0.012839 *
trt2        0.542077  0.051545 10.517 2.57e-08 ***
trt3       -0.224702  0.051545 -4.359 0.000561 ***
trt4       -0.338553  0.051545 -6.568 8.91e-06 ***
trt5       -0.547229  0.051545 -10.617 2.26e-08 ***
trt6        0.562861  0.051545 10.920 1.55e-08 ***
---

```

Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1

Residual standard error: 0.1042 on 15 degrees of freedom

Multiple R-squared: 0.9664, Adjusted R-squared: 0.9396

F-statistic: 35.98 on 12 and 15 DF, p-value: 7.699e-09

```
> model.effects(fit2, "trt")
```

All the usual things such as coefficients, pairwise comparisons, contrasts, etc work.

```

      1      2      3      4      5      6
0.1455303 0.5420770 -0.2247025 -0.3385527 -0.5472291 0.5628608
      7
-0.1399839

```

```
> linear.contrast(fit2, trt, c(1, -1, 0, 0, 0, 0))
```

```

  estimates      se  t-value      p-value  lower-ci  upper-ci
1 -0.3965467 0.07873613 -5.036401 0.0001475959 -0.5643688 -0.2287246

```

```
> sqrt(.010849*(1/4 + (-1)^2/4));sqrt(.010849*(1/3.5 + (-1)^2/3.5))
```

Here we have two attempts to compute the standard error of a pairwise difference. The first is a naive application of the standard formula; this gives the incorrect answer. The second version uses the effective sample size (3.5) and gets the correct answer.

```
[1] 0.0736512
[1] 0.07873645
```

```
> fit3 <- lmer(log(wear) ~ trt + (1|block), data=weardata)
```

If you treat blocks as random and fit the model, then you get interblock recovery.

```
> fit3
```

Note that the standard errors for treatment effects are (slightly) smaller here. That is the result of incorporating the interblock information. However, even with this relatively small block to block variance, the improvement from interblock recovery is small.

Linear mixed model fit by REML [`'lmerMod'`]

Formula: `log(wear) ~ trt + (1 | block)`

Data: `weardata`

REML criterion at convergence: -18.5

Random effects:

```

Groups   Name             Variance Std.Dev.
block    (Intercept)  0.002251 0.04744
Residual                    0.010849 0.10416

```

Number of obs: 28, groups: block, 7

Fixed effects:

| | Estimate | Std. Error | t value |
|-------------|----------|------------|---------|
| (Intercept) | 5.75718 | 0.02663 | 216.21 |
| trt1 | 0.13715 | 0.04964 | 2.76 |
| trt2 | 0.56873 | 0.04964 | 11.46 |
| trt3 | -0.23124 | 0.04964 | -4.66 |
| trt4 | -0.32720 | 0.04964 | -6.59 |
| trt5 | -0.54404 | 0.04964 | -10.96 |
| trt6 | 0.55415 | 0.04964 | 11.16 |

Correlation of Fixed Effects:

| | (Intr) | trt1 | trt2 | trt3 | trt4 | trt5 |
|------|--------|--------|--------|--------|--------|--------|
| trt1 | 0.000 | | | | | |
| trt2 | 0.000 | -0.167 | | | | |
| trt3 | 0.000 | -0.167 | -0.167 | | | |
| trt4 | 0.000 | -0.167 | -0.167 | -0.167 | | |
| trt5 | 0.000 | -0.167 | -0.167 | -0.167 | -0.167 | |
| trt6 | 0.000 | -0.167 | -0.167 | -0.167 | -0.167 | -0.167 |

> **Anova (fit3, test="F")**

The KR anova yields a slightly higher F and slightly higher error df, the result of more information.

Analysis of Deviance Table (Type II Wald F tests with Kenward-Roger df)

Response: log(wear)

| | F | Df | Df.res | Pr(>F) |
|-----|-------|----|--------|---------------|
| trt | 61.09 | 6 | 17.337 | 1.036e-10 *** |

Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1

> **fixef (fit3)**

We can extract the estimates of the fixed effects like this.

| (Intercept) | trt1 | trt2 | trt3 | trt4 | trt5 | trt6 |
|-------------|-----------|-----------|------------|------------|------------|-----------|
| 5.7571762 | 0.1371457 | 0.5687293 | -0.2312409 | -0.3272006 | -0.5440361 | 0.5541537 |

> **linear.contrast (fit3, trt, c(0, 1, -1, 0, 0, 0, 0))**

This should work, but it is failing with an error. Looks like some bug chasing and an update are due.

```
> #  
      OK, we'll do it the hard way for now.  
  
> cfs <- c(0,1,-1,0,0,0,0)  
      These coefficients compare treatments 1 and 2.  
  
> sum(fixef(fit3)*cfs)  
      This is the interblock recovery estimate of the contrast. It differs a bit from the intrablock  
      estimate.  
[1] -0.4315836  
  
> V <- vcov(fit3)  
      V is the variance/covariance matrix of the fixed effects from the model.  
  
> t(cfs)%*%V%*%cfs  
      This matrix product computes the variance of the contrast.  
  
1 x 1 Matrix of class "dgeMatrix"  
      [,1]  
[1,] 0.005750449  
  
> sqrt(.005750449)  
      Take the square root for the standard error of the contrast. Note that it is just a touch less  
      than the SE of the intrablock contrast.  
[1] 0.07583172
```

```
> fit3.mcmc <- lmer.mcmc(fit3, 50000)
```

Let's try the Bayesian approach.

```
> lmer.mcmc.intervals(fit3.mcmc)
```

Curiously, there is little evidence of block to block variability.

| | lower | median | upper | SE |
|-------------|--------------|-------------|-------------|-------------|
| (Intercept) | 5.706598915 | 5.75651753 | 5.80787760 | 0.025473316 |
| trt1 | 0.011344197 | 0.13518825 | 0.25573595 | 0.060744130 |
| trt2 | 0.469056890 | 0.58035595 | 0.69466516 | 0.056392167 |
| trt3 | -0.345104486 | -0.23577003 | -0.12729795 | 0.055242499 |
| trt4 | -0.431603139 | -0.32055717 | -0.20289277 | 0.057299330 |
| trt5 | -0.658420145 | -0.54375852 | -0.43737995 | 0.055216869 |
| trt6 | 0.441788809 | 0.54709335 | 0.65125665 | 0.053101939 |
| block | 0.000000000 | 0.00000000 | 0.01333745 | 0.004718074 |
| sigma2 | 0.006831588 | 0.01283406 | 0.02631961 | 0.004965891 |

```
> m <- fit3.mcmc$mcmcout
```

Get the matrix of Markov chain outputs.

```
> d <- m[, 2]-m[, 3]
```

Treatment 1 and 2 are coefficients 2 and 3, so take their difference in the MCMC samples.

```
> mean(d)
```

Just a touch different from the non-Bayesian estimate.

```
[1] -0.4455276
```

```
> sd(d)
```

As is typical, the Bayesian SE is slightly larger.

```
[1] 0.09033554
```